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CONTENTS

Sr. No.	TITLE & NAME OF THE AUTHOR (S)	Page No.
1.	COUNTRY CHARACTERISTICS AND INFLATION: A PANEL ANALYSIS DR. WILLIAM R. DIPIETRO	1
2.	ROLE OF FINANCIAL MANAGERS IN GLOBAL FINANCIAL CRISIS DR. HAMID SAREMI	4
3.	PATIENT SATISFACTION IN TERTIARY PRIVATE HOSPIATL IN DHAKA: A CASE STUDY ON SQUARE HOSPITAL LTD. SYED HABIB ANWAR PASHA	9
4.	CAPITAL STRUCTURE PATTERNS: A STUDY OF COMPANIES LISTED ON THE COLOMBO STOCK EXCHANGE IN SRI LANKA DR. BALASUNDARAM NIMALATHASAN	16
5.	CORPORATE GOVERNANCE, COMPANY ATTRIBUTES AND VOLUNTARY DISCLOSURES: A STUDY OF NIGERIAN LISTED COMPANIES DR. UMOREN ADEBIMPE & OKOUGBO PEACE	20
6.	CURRENCY FUTURES TRADING IN INDIA	30
7.	IMPACT OF CASA DEPOSIT GROWTH ON THE PROFITABILITY OF NSE LISTED NATIONALIZED BANKS AND NEW GENERATION BANKS IN INDIA - A COMPARATIVE STUDY R. AMUTHAN & DR. A. RAMA CHANDRAN	33
8.	EMERGING NEW MARKET PENAEUS VANNAMEI CULTURE IN INDIA ASLAM CHINARONG & DR B.YAMUNA KRISHNA	38
9.	PRICE DISCOVERY IN THE COMMODITY MARKETS: THE CASE OF FEEDER CATTLE AND LIVE CATTLE MARKETS S. JACKLINE & DR. MALABIKA DEO	42
10.	CUSTOMER RELATIONSHIP MANAGEMENT IN RETAILING WITH SPECIAL REFERNCE TO FAST MOVING CONSUMER GOODS IN ERODE DISTRICT, TAMILNADU, INDIA DR. T. VETRIVEL	47
11.	PRODUCT- THE FIRST 'P' (OF 7P'S) IN INDIAN LIFE INSURANCE SECTOR: AN EMPIRICAL STUDY GANESH DASH & DR. M. BASHEER AHMED KHAN	53
12.	INVESTORS' PERCEPTION TOWARDS THE INFLUENCE OF SPERTEL RISKS ON THE VALUE OF EQUITY SHARES: A STUDY CONDUCTED AT COIMBATORE CITY E. BENNET & DR. M. SELVAM	61
13.	A STUDY OF CONSUMER ATTITUDE TOWARDS CHINESE PRODUCTS (TOYS) IN INDIA WITH SPECIAL REFERENCE TO JALGAON DISTRICT IN MAHARASHTRA PROF. YOGESH D MAHAJAN	66
14.	A STUDY ON FACTORS THAT MOTIVATE IT AND NON-IT SECTOR EMPLOYEES: A COMPARISON DR. S. SARASWATHI	72
15.	A STUDY ON WCM AND PROFITABILITY AFFILIATION DR. AMALENDU BHUNIA & SRI GAUTAM ROY	78
16.	DO GENDER DIFFERENCES IMPACT PROFESSIONAL DEVELOPMENT? DR. VARSHA DIXIT & DR. SUNIL KUMAR	83
17.	EMPLOYEES' PERCEPTION TOWARDS HUMAN RESOURCE PRACTICES IN AIRPORTS AUTHORITY OF INDIA AT CHENNAI DR. PRIYA MANI	87
18.	TECHNICAL ANALYSIS - A PARANORMAL PHENOMENON HARISH GAUTAM	102
19.	SUPPLY AND UTILISATION PATTERN OF AGRICULTURAL CREDIT: A STUDY OF SELECTED CREDIT INSTITUTIONS OF HARYANA DR. SANDEEP CHAHAL	105
20	ADVERTISING THROUGH SOCIAL MEDIA NETWORKS: LET'S CATCH UP WITH THE INTERNET AUDIENCE DR. GAJENDRA SINGH CHAUHAN	112
21	A LITERATURE SURVEY ON EMOTIONAL INTELLIGENCE SHOULD MATTER TO MANAGEMENT YOGESHWER SINGH RANDHAWA & DR. POOJA OHRI	115
22	IDENTIFICATION OF POTENTIAL COMMERCIAL LOCATIONS IN PATNA URBAN AREA AJAY KUMAR & DR. BIJAY KUMAR DAS	117
23	FOREIGN DIRECT INVESTMENT AND ITS IMPACT ON TECHNOLOGY DIFFUSION: SOME ISSUES AND CHALLENGES AHEAD PABITRA KUMAR JENA & RASHI TAGGAR	126
24	AN EMPIRICAL INVESTIGATION INTO THE DETERMINANTS OF FINANCIAL PERFORMANCE OF INDIAN CORPORATE SECTOR: SIZE, GROWTH, LIQUIDITY, PROFITABILITY, DIVIDEND, LEVERAGE BIDYUT JYOTI BHATTACHARJEE	133
25	EMPLOYEE LAY OFF IN MERGER AND ACQUISITION-A CASE STUDY OF AVIATION COMPANIES IN INDIA RAHUL	143
	REQUEST FOR FEEDBACK	146

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ii

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• Schemenner, R.W., Huber, J.C. and Cook, R.L. (1987), "Geographic Differences and the Location of New Manufacturing Facilities," Journal of Urban Economics, Vol. 21, No. 1, pp. 83-104.

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PRICE DISCOVERY IN THE COMMODITY MARKETS: THE CASE OF FEEDER CATTLE AND LIVE CATTLE MARKETS

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ABSTRACT

This paper examines the relationship between the futures market and spot market for the feeder cattle and live cattle markets during the sample period January 2001 through May 2010 and quantifies the price discovery function of commodity futures prices in relation to spot prices of the sample markets. The cointegration tests and Vector Error Correction Models (VECM) employed in the study proved that both the selected futures markets share and provide certain long-run price information to cash markets and they have cointegrated. Both these markets were found to respond favourably to the price discovery mechanism and acted in a similar way.

KEYWORDS

Cointegration, Feeder Cattle and Live Cattle, Price Discovery.

INTRODUCTION

utures markets generally perform two important roles, hedging of risks (in other words called as risk transfer) and price discovery. The efficacy of the hedging function is dependent on the price discovery process or how well new information is reflected in price. Price discovery, or transmission of information into prices, is a crucial function of any markets. Price discovery takes place when order flow from different types of traders is aggregated in a single market, which can be a physical exchange floor or an electronic trading system. This aggregation of trading interests allows for trade prices to correctly represent supply and demand, although market frictions, noise trading, and investor psychology ensure that observed prices are imperfect proxies for the underlying asset values. As markets evolve, it is imperative that the new market structures and trading protocols continue to provide reliable price discovery. In general, futures markets are found to respond faster to new information than spot markets since the transaction cost is lower and the degree of leverage attainable is higher. Whether the spot or the futures market is the center of price discovery in commodity markets has for a long time been discussed in the literature. Stein (1961) showed that futures and spot prices for a given commodity are determined simultaneously. Garbade and Silver (1983) (GS thereafter) developed a model of simultaneous price dynamics in which they established that price discovery takes place in the market with highest number of participants. Their empirical application concludes that "about 75 percent of new information is incorporated first in the future prices." More recently, the price discovery research has focused on microstructure models and on methods to measure it.

SIGNIFICANCE OF THE STUDY AND OBJECTIVES

Price discovery is a concept used frequently, but seldom defined. Thomsen and Foote defined price discovery in 1952 as the process of buyers and sellers arriving at a transaction price for a specific quantity and quality of a commodity or product at a specific time and place. Their definition allows focusing on many interrelated components of the pricing process, and numerous topics may be categorized as price discovery research. Examples include studies of transaction prices and relationships with underlying supply and demand determinants; price relationships and dynamics between and among vertical stages in the marketing channel; spot versus forecasted or futures market prices; price impacts associated with market information, especially public reports; price and product characteristic relationships; spatial and temporal price patterns and dynamics; and price impacts associated with market structure and behaviour changes.

Price discovery research in commodity market has become an increasingly important, because of structural and behavioural changes in storable and non storable commodities, both horizontal and vertical, and the resulting potential price and market information impacts. Structural and behavioural changes in meatpacking and related stages in the livestock-meat subsector have raised questions about price discovery for various species and classes of livestock (Purcell and Rowsell, 1987). Hence, this study attempts to provide empirical evidence on the price discovery mechanisms, taking into consideration the feeder cattle and live cattle.

The objectives so framed for the study were as follows:

- 1. To examine the relationship between the futures market and spot market for the sample commodities;
- 2. To quantify the price discovery function of commodity futures prices in relation to spot prices of feeder cattle and live cattle; and,
- 3. To analyze whether the price information reflects first in futures market or in the spot market.

METHODOLOGY

JOHANSEN'S VECTOR ERROR CORRECTION MODEL (VECM)

Johansen's (1988) Vector Error Correction Model (VECM) was employed to investigate the causal relationship between spot and futures prices. The following steps were followed to estimate Johansen's Vector Error Correction Model (VECM).

Step 1: The stationarity of the data series was evaluated by Augmented Dickey-Fuller (ADF), and Phillips-Perron (PP) tests.

Step 2: Once the series were found integrated in an identical order, then Johansen Multivariate Maximum likelihood cointegration test was employed to investigate the long-run relationship between spot and futures prices and it is presented below.

$$\Delta X_t = \sum_{i=1}^{p-1} \Gamma_i X_{t-i} + \prod X_{t-1} + \varepsilon_t; \ \varepsilon_t = \begin{pmatrix} \varepsilon_{s,t} \\ \varepsilon_{F,t} \end{pmatrix} \approx N(0,\Sigma)$$
(1)

Where $X_t = (S_t F_t)^t$ is the vector of spot and futures prices, Δ denotes the first difference operator; Γ_i and Π are 2×2 coefficient matrices measuring the short-and long-run adjustment of the system to change in X_{ti} and ϵ_t is 2×1 vector of white noise error terms.

Step 3: The test results were quite sensitive to the lag length. Hence, the lag length P was selected on the basis of multivariate generalizations of Akaike's information criteria (AIC).

Step 4: The two likelihood ratio tests were employed to identify the co-integration between the two series. The first statistic λ_{trace} tests whether the number of cointegrating vectors is zero or one, and the other λ_{max} tests whether a single cointegrating equation was sufficient or if two were required. In general, if r cointegrating vector is correct. The following test statistics can be constructed as:

$$\lambda_{tracs} (r) = -T \sum_{i=r+1} ln \left(1 - \frac{\Lambda}{\lambda_{-i}} \right)$$

$$\lambda_{Max} (r, r+1) = -T ln \left(1 - \frac{\Lambda}{\lambda_{-i+1}} \right)$$
(3)

$$\lambda_{Max}\left(r,r+1\right) = -T \ln\left(1 - \frac{\Lambda}{\lambda}_{r+1}\right)$$

Where, n is the number of separate series to be examined, T is the number of usable observations and (λ^{i}) are the estimated Eigen values (also called characteristic roots) obtained from the (i+1) × (i+1) 'cointegrating matrix.'

The first test statistic (λ_{trace}) tests whether the number of distinct cointegrating vectors was less than or equal to r. The second test statistic (λ_{max}) tests the null that the number of cointegrating vectors is r against an r+1. MacKinnon-Haug-Michelis (1999) provide the critical values of these statistics. The rank of Π may be tested using the λ_{max} and λ_{trace} . If rank (Π) =1, then there was single cointegrating vector and Π can be factored as $\Pi = \alpha\beta'$, where α and β' are 2×1 vectors. Using this factorization β 'represents the vector of cointegrating parameters and α is the vector of error correction coefficients measuring the speed of convergence to the long-run steady state.

Step 5: If spot and futures prices were cointegrated, then causality must exist at least in one direction (Granger, 1986). To test the causality, the following vector error correction model (VECM) is estimated by using ordinary least square (OLS) in each equation. p - 1p - 1

$$\Delta S_{t} = \alpha_{s,0} + \sum_{i=1}^{p} \alpha_{s,i} \Delta S_{t-i} + \sum_{i=1}^{p} b_{s,i} \Delta F_{t-i} + \alpha_{s} Z_{t-1} + s_{s,t}$$
(4)
$$\Delta F_{t} = \alpha_{F,0} + \sum_{i=1}^{p-1} \alpha_{F,i} \Delta S_{t-i} + \sum_{i=1}^{p-1} b_{F,i} \Delta F_{t-i} + \alpha_{F} Z_{t-1} + s_{F,t}$$
(5)

where $a_{S,0}$, $a_{F,0}$ are intercept terms; $a_{S,i}$, $b_{S,i}$, $a_{F,i}$, $b_{F,i}$ are the short-run coefficients and $Z_{t-1} = \beta' X_{t-1}$ is the error correction term from equation 1. In terms of the vector error correction model (VECM) of equation 4 & 5, Ft Granger Causes St if some of the b_{s,i}, coefficients, i =1,2,..., p-1 are not zero and α_s , the

error correction coefficient in the equation for spot prices, is significant at conventional levels. Similarly, Sr Granger causes Ft if some of the art coefficients, i =1,2,..., p-1 are not zero and $\alpha_{\rm f}$ is significant at the conventional levels. These hypotheses can be tested by using either t-tests or F-tests on the joint significance of the lagged estimated coefficients. If both St and Ft Granger cause each other, then there is a feedback relationship between the two markets. Therefore, the error correction coefficients, α_s and α_F serve two purposes.

They are (i) to identify the direction of causality between spot and futures prices and (ii) to measure the speed with which deviations from the long-run relationship are corrected by changes in the spot and futures prices.

The vector error correction model (VECM) equation 4 & 5 provides a framework for valid inference in the presence of I (1) variable. Moreover, the Johansen (1988) procedure provides more efficient estimates of the cointegrating relationship than the Engel and Granger (1987) estimator (Gonzalo, 1994). Also, Johansen (1988) tests are shown to be fairly robust to presence of non normality and heteroscedasticity disturbances (Lee and Tse, 1996).

Since the futures prices series and the spot prices series of both feeder and live cattle appeared to be non stationary, causality test had been ignored.

REVIEW OF RELATED LITERATURE

Many previous papers are focused on the price discovery of cross-listed stocks and assets with the same value base. The existence of a price discovery function in futures markets hinges on whether price changes in futures markets lead price changes in cash markets more often than the reverse. Leuthold (1974) investigated the price performance of Live Beef Cattle on the futures market. From the results it was found that Cash cattle prices were found to be more accurate indicators of subsequent cash cattle price conditions than are the futures prices for distant contracts. In other words, evaluation of live beef cattle price relationships revealed that for distant futures, the cash price is a more accurate indicator of future cash price conditions than is the futures price. Also, the futures price becomes less and less efficient both absolutely and relative to the cash price estimates. In other words, the cash price is more stable than the futures price for distant contracts. Oellermann and Farris (1985) investigated lead lag relation between change in futures and spot price for live beef cattle between 1966 and 1982. The futures price led spot price during nearly every sub period analysed. Based on Granger causality test for various sub samples of their data, they conclude that change in live cattle futures price led change in live cattle spot price. They also found that the spot market responded to change in futures price within one trading day. The authors conclude that futures market was the centre of price discovery for live cattle. They suggest that a likely explanation for the results is that the futures market serves as a focal point for information assimilation. They conclude that the cattle futures market contributes towards a more efficient price discovery process in the underlying spot market for live beef cattle.

A slightly different approach was adopted by Koontz et al. (1990) to study price discovery in the livestock market. Using weekly U.S. cash and futures prices from 1973 through 1984, they investigated the spatial nature of the price discovery process. They adopted the procedure proposed by Geweke (1982) to generate causality tests and measures of interaction between major cash markets, and between cash and futures markets. In general, their findings suggested that there was a high degree of interaction between cash and futures prices. They also identified that the pricing relationships changed over time, reflecting changes in the industry which suggests that the price discovery process is dynamic and is influenced by the structure of the underlying markets. Bessler and Covey (1991) studied the futures/cash price relationships for slaughter cattle, a non-storable commodity. They used daily settlement prices for the nearby live cattle futures contract from August 21, 1985 through August 20, 1986, and daily average cash prices (per cwt.) for direct sale of choice 900-1300lbs. slaughter cattle steers in the Texas-Oklahoma market. Their cash series reflected a direct rather than auction sales market for slaughter cattle. Thus, their cash series included sales

throughout the entire five-day business week. For a sample of 261 observations on daily live cattle prices, they obtained mixed results. Within sample fits (conducted on the first 130 data points) indicated that both cash and futures prices were generated by processes not statistically distinguishable from a random walk. Tests for cointegration based on residuals from a static regression (using the same 130 data points) showed marginal support for the cointegration hypothesis between cash and nearby futures prices. No cointegration was discovered between cash prices and more distant contracts. The results are consistent with the suggestion that the greater the temporal spread between futures and cash prices, the greater the degree of independence. Schroeder and Goodwin (1991) studied the price discovery mechanism for livestock in the periods 1979-1986 and 1975- 1989, respectively. Both studies tested the extent of short-run price discovery, and found that information tends to be discovered first in futures markets and then transferred to cash markets. Both studies also adopted other procedures to verify their results in the long run. They used cointegration procedures to verify that daily cash and futures prices for live hogs didn't share a long-run relationship. They found a short-run relationship between cash and futures prices based on Garbade-Silber model, but failed to find a long-run relationship using either Granger-causality or cointegration procedures. Thus as the process of adding to the existing literature, the present study uses the cointegration tests and VECM tests to study the price discovery mechanisms in the feeder cattle and live cattle markets through the years 2001 to 2010.

EMPIRICAL RESULTS

DATA AND SAMPLE PERIODS

The study period of this research was made during 2nd January 2001 to 31st May 2010. The spot and futures prices of feeder cattle and live cattle were obtained from Multi Commodity Exchange of India (MCX).

UNIT ROOT TEST

A necessary condition to carry out a cointegration test in a time series data is that the data have to be non-stationary at the level, but stationary in the differences. Each series of feeder cattle and line cattle prices were first tested for the existence of a unit root by using Augmented Dickey and Fuller (ADF, 1981). The ADF test uses the existence of a unit root as the null hypothesis, that is:

 $H_0: \alpha = 0$

 $H_1: \alpha \neq 0$

TABLE 1: UNIT ROOT TEST RESULTS: AUGMENTED DICKEY AND FULLER (ADF) TEST

Commodities	Null Hypothesis	t-statistic	p-value*	Existence of Unit Root	
Feeder Cattle	Futures prices series has a Unit Root	-1.934793	0.3164	Yes	
	Spot prices series has a Unit Root	-1.888035	0.3382	Yes	
Live Cattle	Futures prices series has a Unit Root	-2.700976	0.0740	Yes	
	Spot prices series has a Unit Root	-1.609250	0.4777	Yes	

Confidence level $\alpha = 0.05$

The results from Augmented Dickey and Fuller (ADF) tests were reported in Table 1. The optimal number of augmenting lags was determined by using Akaike Information Criterion (AIC). The results show that both spot and futures prices of feeder cattle and live cattle were non-stationary (have unit roots).

TABLE 2: UNIT ROOT TEST RESULTS: PHILLIPS-PERRON (PP) TEST						
Commodities	Null Hypothesis	t-statistic	p-value*	Existence of Unit Root		
Feeder Cattle	Futures prices series has a Unit Root	-1.933917	0.3168	Yes		
	Spot prices series has a Unit Root	-1.774476	0.3936	Yes		
Live Cattle	Live Cattle Futures prices series has a Unit Root		0.0917	Yes		
	Spot prices series has a Unit Root	-1.558817	0.5036	Yes		

Confidence level $\alpha = 0.05$

In order to double check the robustness of the results, the Phillips-Perron (PP) test was employed and the results were given in Table 2. As showed in Table 2, the ADF test fails to reject the null hypothesis of the presence of a unit root for the original-prices series.

COINTEGRATION TEST

Given that the spot and futures prices of feeder cattle and live cattle were integrated at the same order, cointegration techniques was used to determine if a long-run relationship exists between the spot and futures prices. Cointegration is an econometric property of time series variables. If two or more series are themselves non-stationary, but a linear combination of them is stationary, then the series are said to be co-integrated. In practice, cointegration is a means for correctly testing those hypotheses concerning the relationship between two variables having unit roots. The Engle–Granger two step methods (Engle & Granger, 1987) and the Johansen trace test

(Johansen, 1988; 1991) are the two main approaches for testing cointegration. Since the spot-futures prices of feeder cattle and live cattle were non-stationary, in this study, the Johansen trace test was adapted to test the presence of the long-run equilibrium relationship in the hourly prices and trading volume series in pairs. As stated by Johansen (1988), the likelihood ratio-test statistic for the hypothesis of the at most r co-integrated relationship and the at least m = n - r common trend is given by:

$$\lambda_{trace}(r) = -T \sum_{i=r+1}^{n} \ln \left(1 - \frac{\Lambda}{\lambda}_{i}\right)$$

Where T is the sample size, and ² represents Eigen values of the squared canonical correlation between the two residual vectors from the level regressions. The results of the Johansen trace test were reported in Table 3.

TABLE 3: JOHANSEN COINTEGRATION TEST RESOLTS							
Commodities	Hypothesized No. of	Eigenvalue	Trace Statistic	0.05 critical	Max-Eigen	0.05 critical	p-value [*]
	CE (s)			value	Statistic	value	
Feeder Cattle	No cointegration [*]	0.061591	146.4872	15.49471	142.9048	14.26460	0.0001
	Atmost one	0.001592	3.582444	3.841466	3.582444	3.841466	0.0584
	Cointegration						
Live Cattle	No cointegration [*]	0.010283	25.50765	15.49471	23.11238	14.26460	0.0011
	Atmost one	0.001071	2.395265	3.841466	2.395265	3.841466	0.1217
	Cointegration						

Trace test indicates 1 cointegrating eqn(s) at the 0.05 level

* denotes rejection of the hypothesis at the 0.05 level

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As seen in Table 3, the rejection of the "null hypothesis of no cointegration" indicates that the spot and futures price series of the feeder cattle and live cattle have a long-run equilibrium relationship.

VECTOR ERROR CORRECTION MODEL

A vector error correction model (VECM; Engle & Granger, 1987) can lead to a better understanding of the nature of any non-stationarity among the different component series and can also improve the longer-term forecasting over an unconstrained model. Because the spot-futures prices of feeder cattle and live cattle were co-integrated, the VECM was used to investigate the lead-lag relationship of spot and futures prices. Assuming that St is the spot prices series, and Ft denotes the futures prices series, and then the VECM in this study can be illustrated as follows: m = 1

$$\Delta S_{c} = \alpha_{s,0} + \sum_{i=1}^{p-1} \alpha_{s,i} \Delta S_{c-i} + \sum_{i=1}^{p-1} b_{s,i} \Delta F_{c-i} + \alpha_{s} Z_{c-1} + \varepsilon_{s,c}$$
$$\Delta F_{c} = \alpha_{F,0} + \sum_{i=1}^{p-1} \alpha_{F,i} \Delta S_{c-i} + \sum_{i=1}^{p-1} b_{F,i} \Delta F_{c-i} + \alpha_{F} Z_{c-1} + \varepsilon_{F,c}$$

where Δ is the differencing operator, p, q, m, and n are the orders of lag which were determined by Akaike's information criterion evidence, and $\varepsilon_{s,t}$ and EF, t were white noises.

TABLE 4: VECTOR ERROR CORRECTION ESTIMATES RESULTS

Feeder Cattle			Live Cattle		
Cointegrating Eq:	CointEq1		Cointegrating Eq:	CointEq1	
FCFUTURES(-1)	1.000000		LCFUTURES(-1)	1.000000	
FCSPOT(-1)	-0.852299		LCSPOT(-1)	-0.578245	
	(0.02389)			(0.08764)	
	[-35.6824]			[-6.59774]	
С	2.820009		С	-18.81324	
Error Correction:	D(FCFUTURE)	D(FCSPOT)	Error Correction:	D(LCFUTURES)	D(LCSPOT)
CointEq1	-0.059775	-0.000866	CointEq1	-0.014528	0.005952
	(0.00533)	(0.00564)		(0.00344)	(0.00298)
	[-11.2225]	[-0.15360]		[-4.21917]	[2.00026]
D(FCFUTURES(-1))	0.032798	-0. <mark>041151</mark>	D(LCFUTURES(-1))	0.051287	-0.020475
	(0.02061)	(0.02180)		(0.02114)	(0.01827)
	[1.59174]	[-1 <mark>.88</mark> 759]		[2.42609]	[-1.12086]
D(FCFUTURES(-2))	0.022158	-0.002644	D(LCFUTURES(-2))	0.053896	0.017166
	(0.02063)	(0.02183)		(0.02115)	(0.01828)
	[1.07390]	[-0.12111]		[2.54771]	[0.93906]
D(FCSPOT(-1))	-0.058020	0.074044	D(LCSPOT(-1))	-0.028108	0.045975
	(0.02045)	(0.02164)		(0.02448)	(0.02116)
	[-2.83671]	[3.42158]		[-1.14810]	[2.17316]
D(FCSPOT(-2))	-0.049161	-0.011974	D(LCSPOT(-2))	-0.003819	0.028366
	(0.02054)	(0.02173)		(0.02450)	(0.02117)
	[-2.39389]	[-0.55111]		[-0.15591]	[1.34000]
С	0.001506	0.004093	С	0.003902	0.005906
	(0.02030)	(0.02148)		(0.02042)	(0.01765)
	[0.07421]	[0.19059]		[0.19108]	[0.33466]
R-squared	0.054089	0.006771	R-squared	0.012595	0.005470
Adj. R-squared	0.051982	0.004558	Adj. R-squared	0.010383	0.003242

 ${m S}_{{m t}^{-i}}$ means a deviation from the long-run co-integrated equilibrium in the last period. This As stated by Engle and Granger (1987), the error correction term

model interprets that the change in the spot-prices series (S_t) and the futures-prices series (F_t) was due to short-run effects from past $\frac{\Delta}{S_t}$ and $\frac{\Delta}{F_t}$, and to the adjustment of the long-run equilibrium. The final results of the vector-error correction estimates were listed in Table 4. As seen in Table 4, the rejection of the null hypothesis of no VECM indicates that the spot and futures prices series of feeder cattle and live cattle have a cointegrated relationship.

CONCLUSIONS AND IMPLICATIONS

In this paper, an attempt was made to investigate the price discovery mechanisms in the feeder cattle and live cattle markets. Both these markets were found to respond favourably to the mechanism and act in a similar way. The futures and spot prices appeared to be non stationary for these commodities. The evidence obtained supports the fact that the forward pricing role may serve price discovery on commodity markets (Black, 1976; Peck, 1985). Consistent with this finding, in this study also the selected futures markets share and provide certain long-run price information to cash markets for both non storable commodities during the sample periods. The findings disagree somewhat with the prevalent suspicion that a price discovery function would not work at all for non storable commodities because of the lack of storage. However, it also cautions against the naive use of futures prices as expected cash prices for most livestock commodities. In sum it can be concluded that both of these markets share long run equilibrium and they have cointegrated relationships.

The study can be useful to the investors, producers and academicians who are very keen in observing the trend of these markets. Since the interests of the investors on the non storable markets are comparatively low when compared to the agricultural products, research contribution to the price discovery mechanisms on these markets can help in extending its market boundaries. At the same time, it is suggested that in future, the study can still be extended as

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the comparison between some of the storable assets and the non storable assets so that a clear understanding on whether asset storability impacts price discovery can be empirically proven.

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